

# Gaussian Processes for Inference in Biological Interaction Networks

**Neil Lawrence**

School of Computer Science

University of Manchester

Joint work with **Magnus Rattray** and **Guido Sanguinetti**

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# Outline

- 1 Application
  - Methodology & Application Overview
  - Covariance functions
  - Regression with Gaussian Processes
- 2 Latent Functions
  - Toy Problem
  - Biological Problem
- 3 Non-linear Response Model
  - Linear Response with MLP Kernel
  - Non-linear Responses

# Online Resources

All source code and slides are available online

- This talk available from home page (see talks link on side).
- Scripts available in the 'gpsim' toolbox
  - <http://www.cs.man.ac.uk/~neill/gpsim/>.
- MATLAB commands used for examples given in typewriter font.

# Framework

## Latent functions

- Many interaction networks have *latent functions*.
- Assume a Gaussian process (GP) prior distribution for the latent function.
  - Gaussian processes (GPs) are probabilistic models for functions. O'Hagan [1978, 1992], Rasmussen and Williams [2006]
- Our Approach
  - 1 Take a differential equation model for the system.
  - 2 Derive GP covariance jointly for observed and latent functions.
  - 3 Maximise likelihood with respect to parameters (mostly physically meaningful).

# This Talk

## Transcription Network

- Introduce Gaussian Processes for dealing with *latent functions* in transcription networks.
- Show how in a linear response model the latent function can be dealt with *analytically*.
- Discuss extensions to systems with non-linear responses.

# Linear Response Model

## p53 Inference [Barenco et al., 2006]

- Data consists of  $T$  measurements of mRNA expression level for  $N$  different genes.
- We relate gene expression,  $x_j(t)$ , to TFC,  $f(t)$ , by

$$\frac{dx_j(t)}{dt} = B_j + S_j f(t) - D_j x_j(t). \quad (1)$$

$B_j$  basal transcription rate of gene  $j$ ,

$S_j$  is sensitivity of gene  $j$

$D_j$  is the decay rate of the mRNA.

- Dependence of mRNA transcription rate on TF is linear.

# Linear Response Solution

## Solve for TFC

- The equation given in (3) can be solved to recover

$$x_j(t) = \frac{B_j}{D_j} + S_j \exp(-D_j t) \int_0^t f(u) \exp(D_j u) du. \quad (2)$$

- If we model  $f(t)$  as a GP then as (2) only involves *linear operations*  $x_j(t)$  is also a GP.

# Gaussian Processes

## GP Advantages

- GPs allow for inference of continuous profiles, accounting naturally for temporal structure.
  - GPs allow joint estimation of a mRNA concentration and production rates (derivative observations).
  - GPs deal consistently with the uncertainty inherent in the measurements.
  - GPs outstrip MCMC for computational efficiency.

**Note:** GPs have previously been proposed for solving differential equations [Graepel, 2003] and dynamical systems [Murray-Smith and Pearlmuter].



# Defining a Distribution over Functions

## Gaussian Process

- What is meant by a distribution over functions?
- Functions are infinite dimensional objects:
  - Defining a distribution over functions seems non-sensical.

## Gaussian Distribution

- Start with a standard Gaussian distribution.
- Consider the distribution over a fixed number of instantiations of the function.

# Gaussian Distribution

## Zero mean Gaussian distribution

- A multi-variate Gaussian distribution is defined by a mean and a covariance matrix.

$$N(\mathbf{f}|\mu, \mathbf{K}) = \frac{1}{(2\pi)^{\frac{N}{2}} |\mathbf{K}|^{\frac{1}{2}}} \exp\left(-\frac{(\mathbf{f} - \mu)^T \mathbf{K}^{-1} (\mathbf{f} - \mu)}{2}\right).$$

- We will consider the special case where the mean is zero,

$$N(\mathbf{f}|\mathbf{0}, \mathbf{K}) = \frac{1}{(2\pi)^{\frac{N}{2}} |\mathbf{K}|^{\frac{1}{2}}} \exp\left(-\frac{\mathbf{f}^T \mathbf{K}^{-1} \mathbf{f}}{2}\right).$$

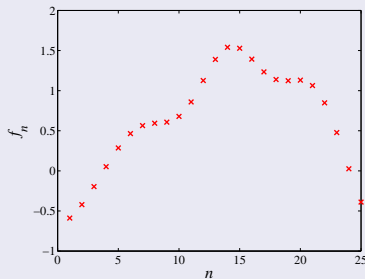
# Sampling a Function

## Multi-variate Gaussians

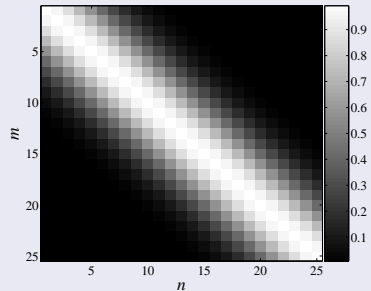
- We will consider a Gaussian with a particular structure of covariance matrix.
- Generate a single sample from this 25 dimensional Gaussian distribution,  $\mathbf{f} = [f_1, f_2 \dots f_{25}]$ .
- We will plot these points against their index.

# Gaussian Distribution Sample

demGPSample



(a)



(b)

**Figure:** (a) 25 instantiations of a function,  $f_n$ , (b) greyscale covariance matrix.

# Covariance Function

## The covariance matrix

- Covariance matrix shows correlation between points  $f_m$  and  $f_n$  if  $n$  is near to  $m$ .
- Less correlation if  $n$  is distant from  $m$ .
- Our ordering of points means that the *function appears smooth*.
- Let's focus on the joint distribution of two points from the 25.

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# Prediction of $f_2$ from $f_1$

```
demGPCov2D([1 2])
```

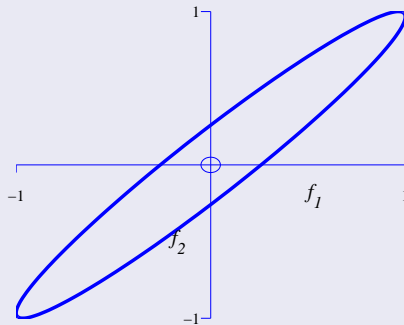


Figure: Covariance for  $\begin{bmatrix} f_1 \\ f_2 \end{bmatrix}$  is  $\mathbf{K}_{12} = \begin{bmatrix} 1 & 0.966 \\ 0.966 & 1 \end{bmatrix}$ .

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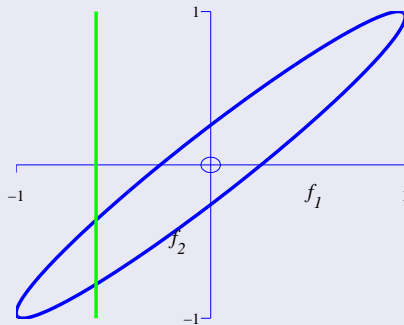


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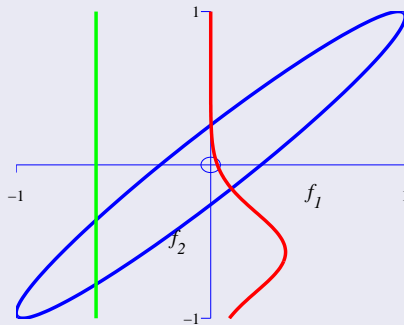


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# Prediction of $f_5$ from $f_1$

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demGPCov2D([1 5])
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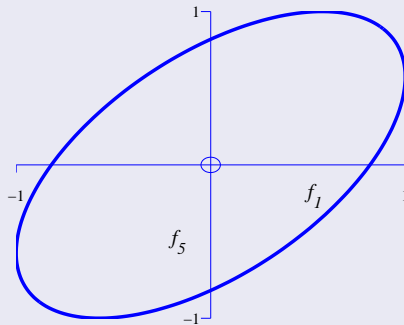


Figure: Covariance for  $\begin{bmatrix} f_1 \\ f_5 \end{bmatrix}$  is  $\mathbf{K}_{15} = \begin{bmatrix} 1 & 0.574 \\ 0.574 & 1 \end{bmatrix}$ .

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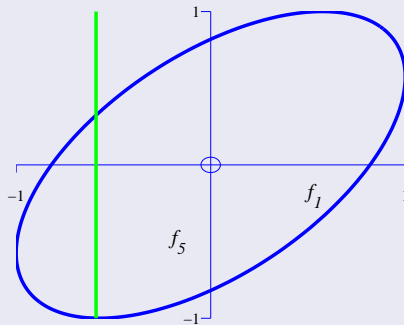


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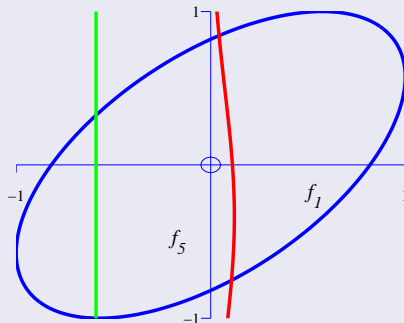


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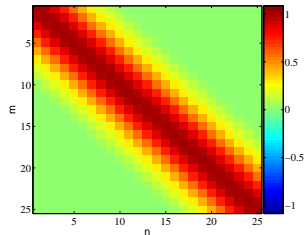
# Covariance Functions

## Visualisation of RBF Covariance

### RBF Kernel Function

$$k(t, t') = \alpha \exp\left(-\frac{(t - t')^2}{2l^2}\right)$$

- Covariance matrix is built using the time *inputs* to the function,  $t$ .
- For the example above it was based on Euclidean distance.
- The covariance function is also known as a kernel.



# Covariance Samples

demCovFuncSample – sample from the prior

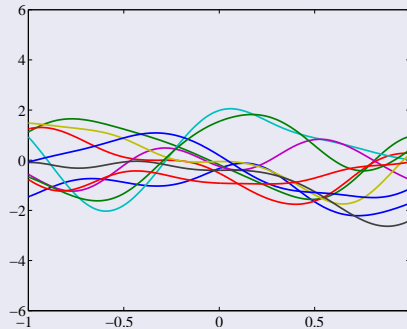


Figure: RBF kernel with  $l = 0.3$ ,  $\alpha = 1$



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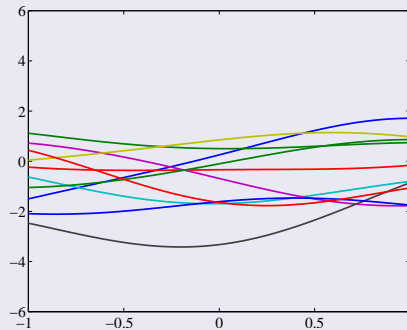


Figure: RBF kernel with  $l = 1$ ,  $\alpha = 1$

# Covariance Samples

demCovFuncSample – sample from the prior

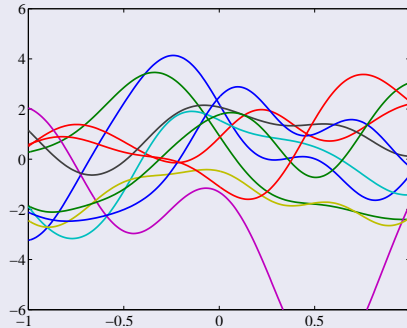


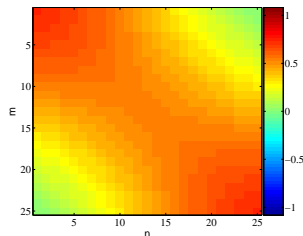
Figure: RBF kernel with  $l = 0.3$ ,  $\alpha = 4$

# Different Covariance Functions

## MLP Kernel Function

$$k(t, t') = \alpha \sin^{-1} \left( \frac{wtt' + b}{\sqrt{wt^2 + b + 1} \sqrt{wt'^2 + b + 1}} \right)$$

- A non-stationary covariance matrix Williams [1997].
- Derived from a multi-layer perceptron (MLP).



# Covariance Samples

demCovFuncSample — samples from the prior

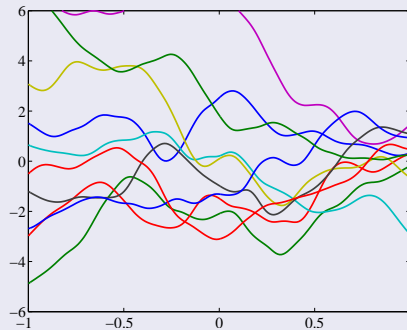


Figure: MLP kernel with  $\alpha = 8$ ,  $w = 100$  and  $b = 100$

# Covariance Samples

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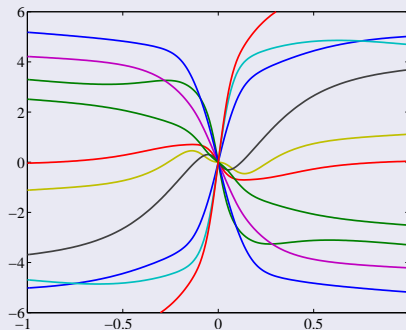


Figure: MLP kernel with  $\alpha = 8$ ,  $b = 0$  and  $w = 100$

# Prior to Posterior

## Prediction with GPs

- GPs provide a probabilistic prior over functions.
- By combining with data we get a *posterior* over functions.
- This is obtained through combining a covariance function with data.
- Toy Example: regression with GPs.

# Gaussian Process Regression

demRegression

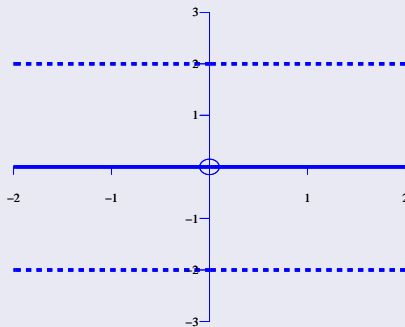


Figure: Going from prior to posterior with data.

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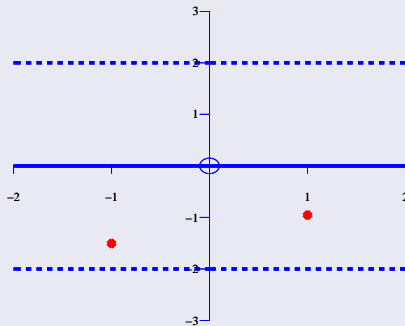


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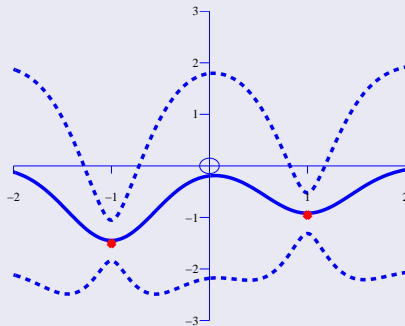


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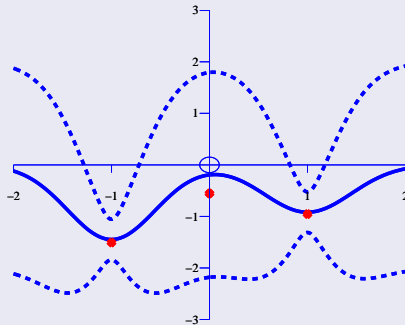


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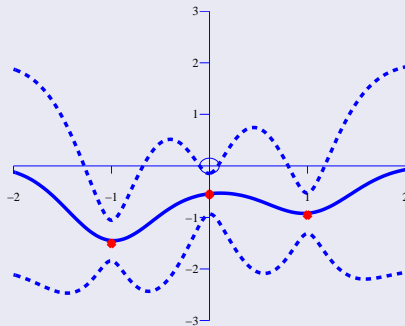


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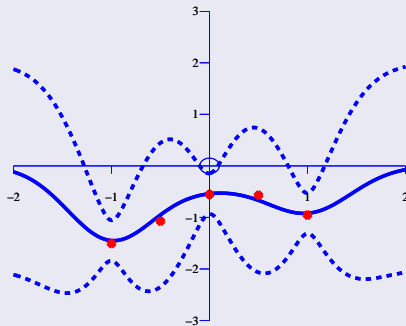


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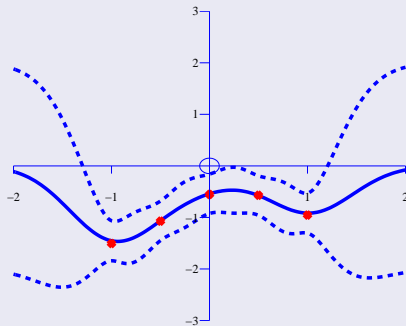


Figure: Going from prior to posterior with data.

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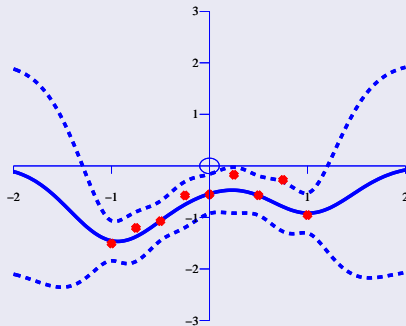


Figure: Going from prior to posterior with data.

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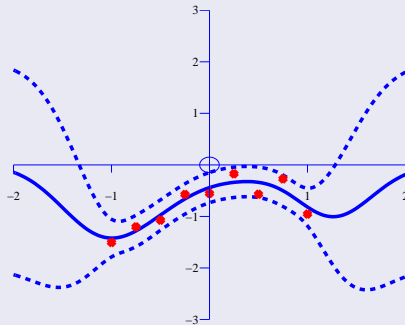


Figure: Going from prior to posterior with data.

# Linear Response Model

## p53 Inference [Barenco et al., 2006]

- Recall Barenco et al.'s linear response model.

$$\frac{dx_j(t)}{dt} = B_j + S_j f(t) - D_j x_j(t). \quad (3)$$

$B_j$  basal transcription rate of gene  $j$ ,

$S_j$  is sensitivity of gene  $j$

$D_j$  is the decay rate of the mRNA.

- We will place a *prior distribution* over the *latent function*.



# Covariance of Latent Function

## Prior Distribution for TFC

- We assume that the TF concentration is a Gaussian Process.
- We will assume an RBF covariance function

$$p(\mathbf{f}) = N(\mathbf{f} | \mathbf{0}, \mathbf{K}) \quad k(t, t') = \exp\left(-\frac{(t - t')^2}{2l^2}\right).$$

# Computation of Joint Covariance

## Covariance Function Computation

- We rewrite solution of differential equation as

$$x_j(t) = \frac{B_j}{D_j} + L_j[f](t)$$

where

$$L_j[f](t) = S_j \exp(-D_j t) \int_0^t f(u) \exp(D_j u) du \quad (4)$$

is a linear operator.

# Induced Covariance

## Gene's Covariance

- The new covariance function is then given by

$$\text{cov} \left( L_j [f] (t), L_k [f] (t') \right) = L_j \otimes L_k [k_{ff}] (t, t') .$$

more explicitly

$$\begin{aligned} k_{x_j x_k} (t, t') &= S_j S_k \exp (-D_j t - D_k t') \int_0^t \exp (D_j u) \\ &\quad \times \int_0^{t'} \exp (D_k u') k_{ff} (u, u') du' du . \end{aligned}$$

- With RBF covariance these integrals are tractable.

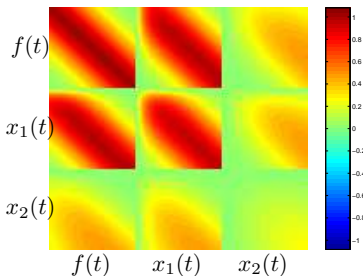
# Covariance for Transcription Model

## RBF Kernel function for $f(t)$

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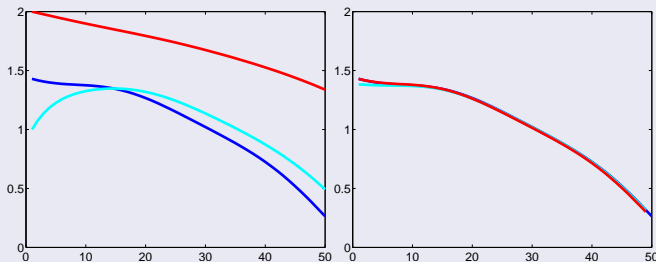
- Joint distribution for  $x_1(t)$ ,  $x_2(t)$  and  $f(t)$ .
- Here:

$D_1$	$S_1$	$D_2$	$S_2$
5	5	0.5	0.5



# Joint Sampling of $x(t)$ and $f(t)$

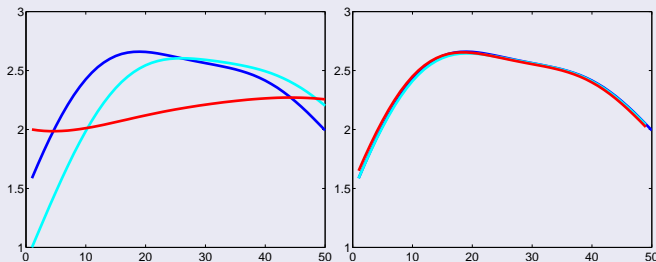
gpsimTest



**Figure:** *Left:* joint samples from the transcription covariance, *blue:*  $f(t)$ , *cyan:*  $x_1(t)$  and *red:*  $x_2(t)$ . *Right:* numerical solution for  $f(t)$  of the differential equation from  $x_1(t)$  and  $x_2(t)$  (blue and cyan). True  $f(t)$  included for comparison.

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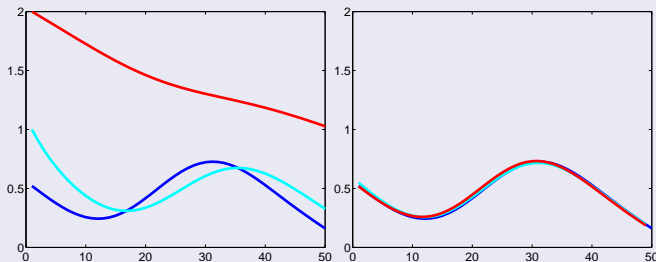
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# Noise Corruption

## Estimate Underlying Noise

- Allow the mRNA abundance of each gene at each time point to be corrupted by noise, for observations at  $t_i$  for  $i = 1, \dots, T$ ,

$$y_j(t_i) = x_j(t_i) + \epsilon_j(t_i) \quad (5)$$

with  $\epsilon_j(t_i) \sim \mathcal{N}(0, \sigma_{ji}^2)$ .

- Estimate noise level using probe-level processing techniques of Affymetrix microarrays (e.g. mmgMOS, [Liu et al., 2005]).
- The covariance of the noisy process is then  $K_{yy} = \Sigma + K_{xx}$ , with  $\Sigma = \text{diag}(\sigma_{11}^2, \dots, \sigma_{1T}^2, \dots, \sigma_{N1}^2, \dots, \sigma_{NT}^2)$ .



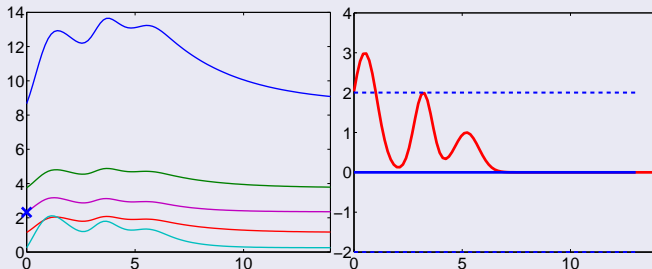
# Artificial Data

## Toy Problem

- Results from an artificial data set.
- We used a 'known TFC' and derived six 'mRNA profiles'.
  - Known TFC composed of three Gaussian basis functions.
  - mRNA profiles derived analytically.
- Fourteen subsamples were taken and corrupted by noise.
- This 'data' was then used to:
  - Learn decays, sensitivities and basal transcription rates.
  - Infer a posterior distribution over the missing TFC.

# Artificial Data Results

## demToyProblem1

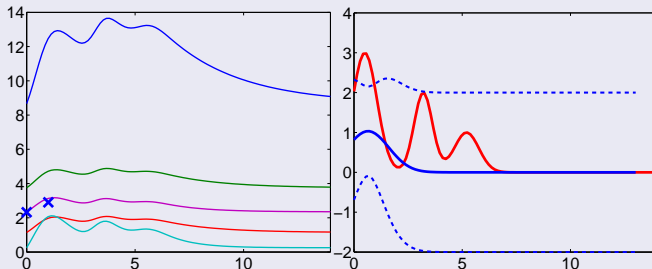


**Figure:** *Left:* The TFC,  $f(t)$ , which drives the system. *Middle:* Five gene mRNA concentration profiles each obtained by using different parameter sets  $\{B_i, S_i, D_i\}_{i=1}^5$  (lines) along with noise corrupted 'data'. *Right:* The inferred TFC (with error bars).

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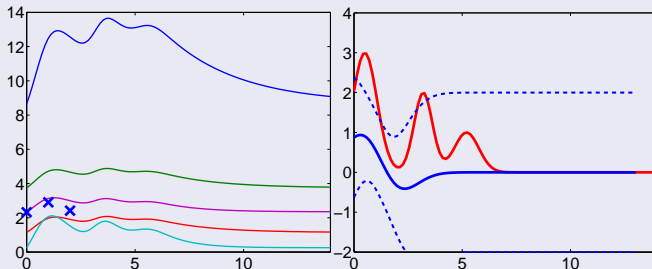


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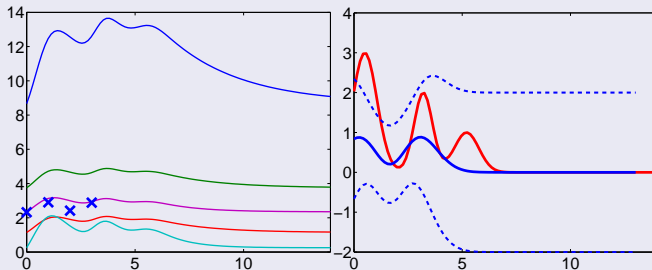


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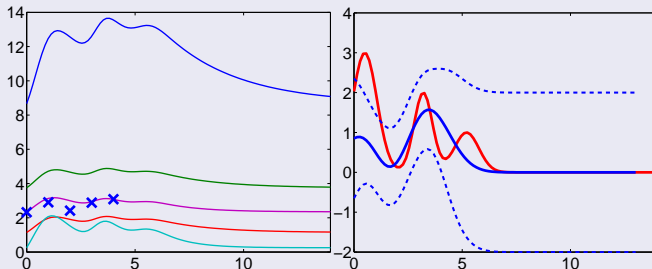


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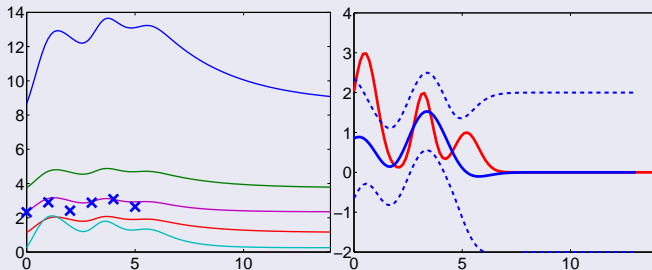


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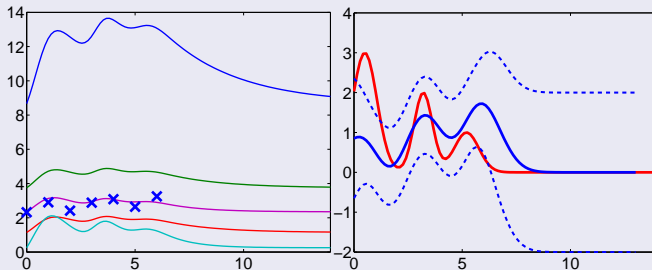


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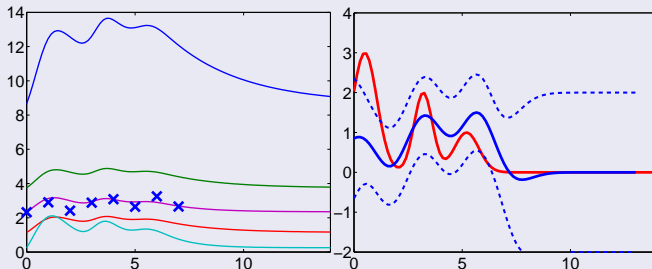
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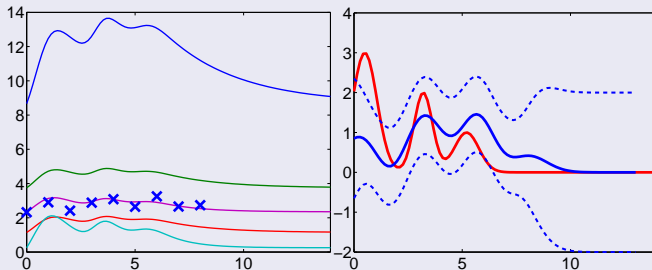


**Figure:** *Left:* The TFC,  $f(t)$ , which drives the system. *Middle:* Five gene mRNA concentration profiles each obtained by using different parameter sets  $\{B_i, S_i, D_i\}_{i=1}^5$  (lines) along with noise corrupted 'data'. *Right:* The inferred TFC (with error bars).

## demToyProblem1

# Artificial Data Results

## demToyProblem1

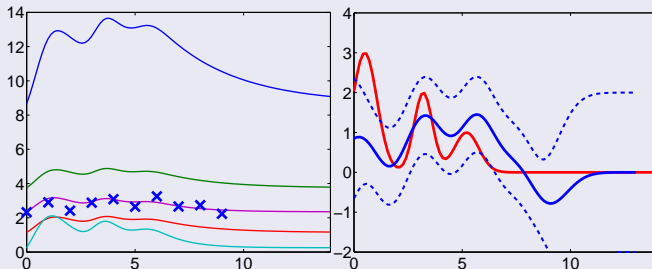


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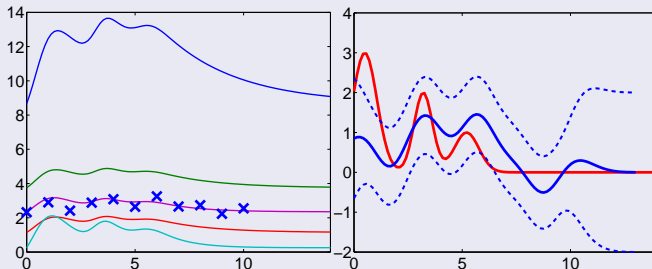


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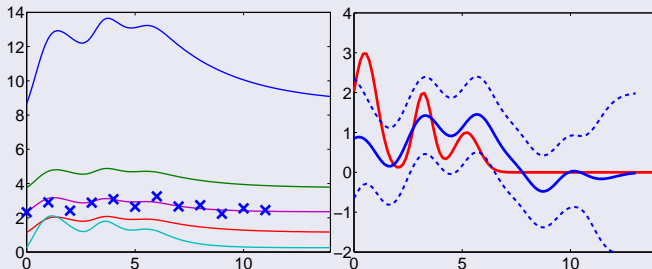


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# Artificial Data Results

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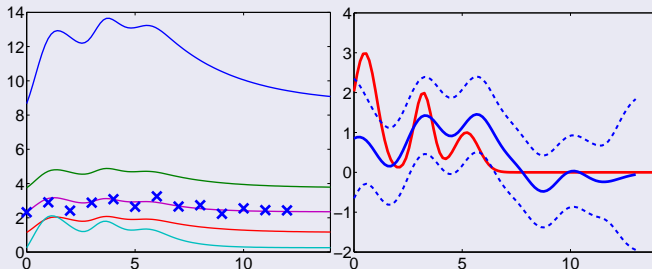


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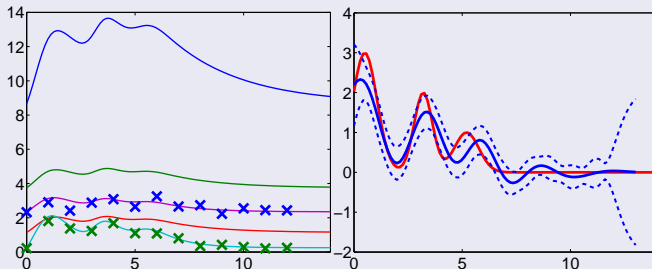


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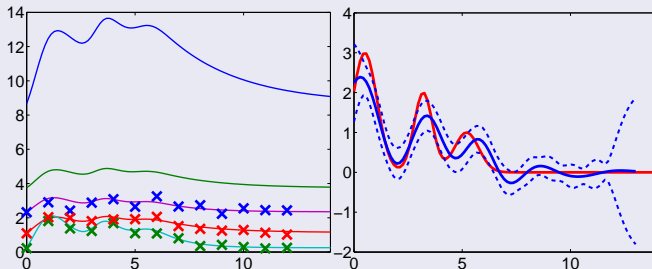


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## demToyProblem1

# Artificial Data Results

## demToyProblem1



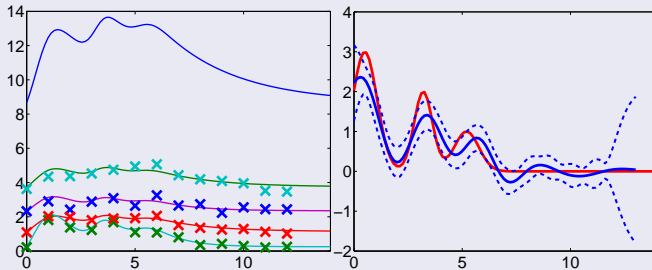
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## demToyProblem1



# Artificial Data Results

## demToyProblem1

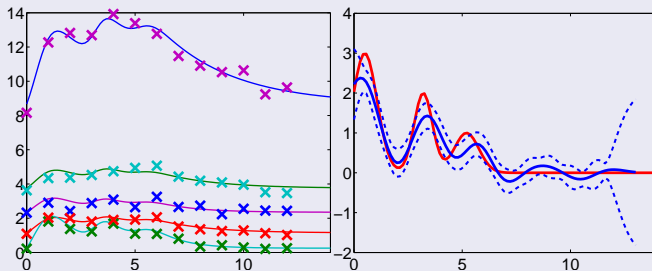


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## demToyProblem1



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## demToyProblem1

# Results

## Linear System

- Recently published biological data set studied using linear response model by Barenco et al. [2006].
- Study focused on the tumour suppressor protein p53.
- mRNA abundance measured for five targets: *DDB2*, *p21*, *SESN1/hPA26*, *BIK* and *TNFRSF10b*.
- Quadratic interpolation for the mRNA production rates to obtain gradients.
- They used MCMC sampling to obtain estimates of the model parameters  $B_j$ ,  $S_j$ ,  $D_j$  and  $f(t)$ .

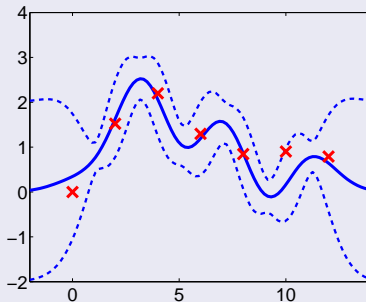
# Linear response analysis

## Experimental Setup

- We analysed data using the linear response model.
- Raw data was processed using the mmgMOS model of Liu et al. [2005] which provides variance as well as expression level.
- We present posterior distribution over TFCs.
- Results of inference on the values of the hyperparameters  $B_j$ ,  $S_j$  and  $D_j$ .
  - Samples from the posterior distribution were obtained using Hybrid Monte Carlo (see e.g. Neal, 1996).

# Linear Response Results

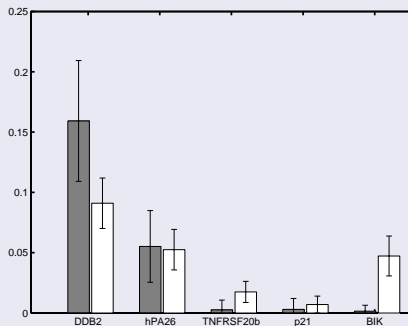
demBarenco1



**Figure:** Predicted protein concentration for p53. Solid line is mean, dashed lines 95% credibility intervals. The prediction of [Barenco et al., 2006] was pointwise and is shown as crosses.

# Results — Transcription Rates

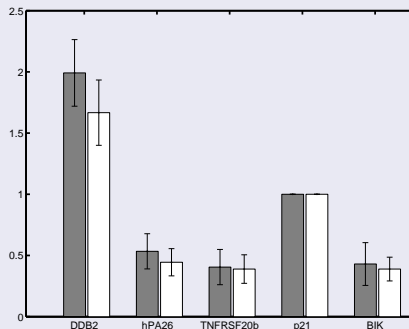
## Estimation of Equation Parameters demBarenco1



**Figure:** Basal transcription rates. Our results (black) compared with Barenco et al. [2006] (white).

# Results — Transcription Rates

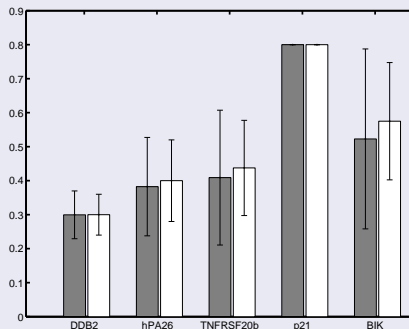
## Estimation of Equation Parameters demBarenco1



**Figure:** Sensitivities. Our results (black) compared with Barenco et al. [2006] (white).

# Results — Transcription Rates

## Estimation of Equation Parameters demBarenco1



**Figure:** Decays. Our results (black) compared with Barenco et al. [2006] (white).



# Linear Response Discussion

## GP Results

- Note oscillatory behaviour, possible artifact of RBF covariance Rasmussen and Williams [see page 123 in 2006].
- Results are in good accordance with the results obtained by Barenco et al..
- Differences in estimates of the basal transcription rates probably due to:
  - different methods used for probe-level processing of the microarray data.
  - Our failure to constrain  $f(0) = 0$ .
- Our results take about 13 minutes to produce Barenco et al. required 10 million iterations of Monte Carlo.

# Non-linear Response Model

## More Realistic Response

- All the quantities in equation (3) are positive, but direct samples from a GP will not be.
- Linear models don't account for saturation.
- *Solution:* model response using a positive nonlinear function.

# Formalism

## Non-linear Response

- Introduce a non-linearity  $g(\cdot)$  parameterised by  $\theta_j$

$$\frac{dx_j}{dt} = B_j + g(f(t), \theta_j) - D_j x_j$$
$$x_j(t) = \frac{B_j}{D_j} + \exp(-D_j t) \int_0^t du g(f(u), \theta_j) \exp(D_j u) .$$

- The induced distribution of  $x_j(t)$  is no longer a GP.
- Derive the functional gradient and learn a MAP solution for  $f(t)$ .
- Also compute Hessian so we can approximate the marginal likelihood.

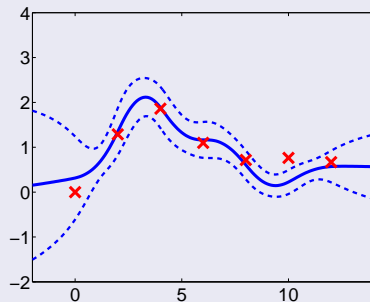
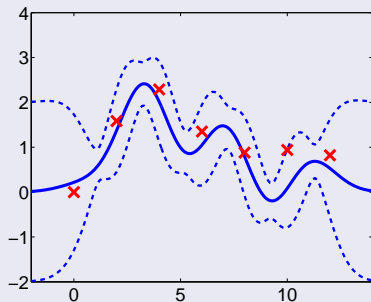
## Example: linear response

### Using non-RBF kernels

- Start by taking  $g(\cdot)$  to be linear.
- Provides 'sanity check' and allows arbitrary covariance functions.
- Avoids double numerical integral that would normally be required.

# Response Results

demBarencoMap1, demBarencoMap2



**Figure:** *Left:* RBF prior on  $f$  (log likelihood -101.4); *Right:* MLP prior on  $f$  (log likelihood -105.6).

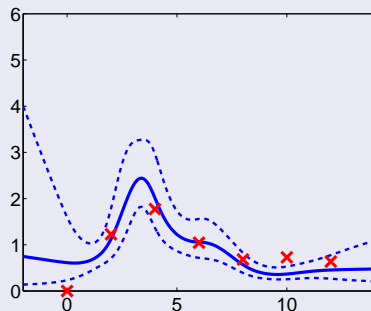
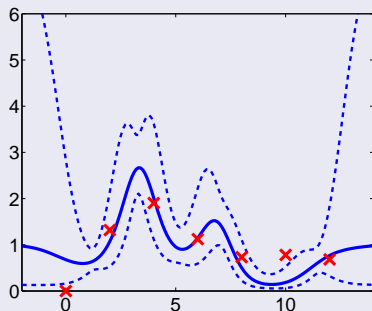
# Non-linear response analysis

## Non-linear responses

- Exponential response model (constrains protein concentrations positive).
- $\log(1 + \exp(f))$  response model.
- $\frac{3}{1 + \exp(-f)}$
- Inferred MAP solutions for the latent function  $f$  are plotted below.

# $\exp(\cdot)$ Response Results

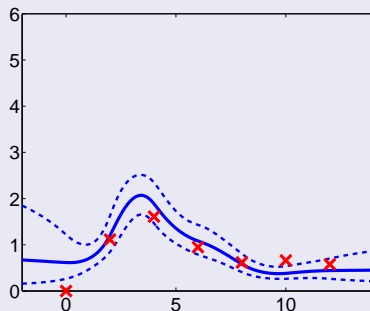
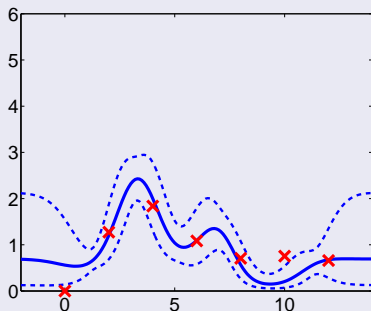
demBarencoMap3, demBarencoMap4



**Figure:** *Left:* shows results of using a squared exponential prior covariance on  $f$  (log likelihood -100.6); *Right:* shows results of using an MLP prior covariance on  $f$  (log likelihood -106.4).

# $\log(1 + \exp(f))$ Response Results

demBarencoMap5, demBarencoMap6

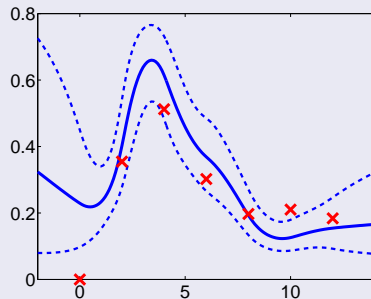
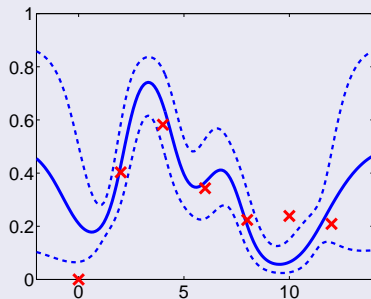


**Figure:** *Left:* shows results of using a squared exponential prior covariance on  $f$  (log likelihood -100.9); *Right:* shows results of using an MLP prior covariance on  $f$  (log likelihood -110.0).



$\frac{3}{1+\exp(-f)}$  Response Results

demBarencoMap7, demBarencoMap8



**Figure:** *Left:* shows results of using a squared exponential prior covariance on  $f$  (log likelihood -104.1); *Right:* shows results of using an MLP prior covariance on  $f$  (log likelihood -111.2).

# Discussion

- We have described how GPs can be used in modelling dynamics of a simple regulatory network motif.
- Our approach has advantages over standard parametric approaches:
  - there is no need to restrict the inference to the observed time points, the temporal continuity of the inferred functions is accounted for naturally.
  - GPs allow us to handle uncertainty in a natural way.
  - MCMC parameter estimation in a discretised model can be computationally expensive. Parameter estimation can be achieved easily in our framework by type II maximum likelihood or by using efficient hybrid Monte Carlo sampling techniques
- All code on-line  
<http://www.cs.man.ac.uk/~neill/gpsim/>.

# Future Directions

## What Next?

- This is still a very simple modelling situation.
  - We are ignoring transcriptional delays.
  - Here we have single transcription factor: our ultimate goal is to describe regulatory pathways with more genes.
  - All these issues can be dealt with in the general framework we have described.
  - Need to overcome the greater computational difficulties.

# Acknowledgements

## Data and Support

We thank Martino Barenco for useful discussions and for providing the data. We gratefully acknowledge support from BBSRC Grant No BBS/B/0076X “Improved processing of microarray data with probabilistic models”.

# Covariance Result

## Covariance Result

$$k_{x_j \times_k}(t, t') = S_j S_k \frac{\sqrt{\pi}}{2} [h_{kj}(t', t) + h_{jk}(t, t')]$$

where

$$\begin{aligned} h_{kj}(t', t) = & \frac{\exp(\gamma_k)^2}{D_j + D_k} \\ & \times \left\{ \exp[-D_k(t' - t)] \left[ \operatorname{erf}\left(\frac{t' - t}{l} - \gamma_k\right) + \operatorname{erf}\left(\frac{t}{l} + \gamma_k\right) \right] \right. \\ & \left. - \exp[-(D_k t' + D_j)] \left[ \operatorname{erf}\left(\frac{t'}{l} - \gamma_k\right) + \operatorname{erf}(\gamma_k) \right] \right\}. \end{aligned}$$

Here  $\gamma_k = \frac{D_k l}{2}$ .

# Cross Covariance

## Correlation of $x_j(t)$ and $f(t')$

- Need the “cross-covariance” terms between  $x_j(t)$  and  $f(t')$ , which is obtained as

$$k_{x_j f}(t, t') = S_j \exp(-D_j t) \int_0^t \exp(D_j u) k_{ff}(u, t') du. \quad (6)$$

- For RBF we have

$$k_{x_j f}(t', t) = \frac{\sqrt{\pi} S_j e^{2\gamma_j}}{2} \exp[-D_j(t' - t)] \left[ \operatorname{erf}\left(\frac{t' - t}{l} - \gamma_j\right) + \operatorname{erf}\left(\frac{t}{l} + \gamma_j\right) \right].$$

# Posterior for $f$

## Prediction for TFC

- Standard Gaussian process regression techniques [see e.g. Rasmussen and Williams, 2006] yield

$$\begin{aligned}\langle f \rangle_{\text{post}} &= K_{f\mathbf{x}} K_{\mathbf{xx}}^{-1} \mathbf{x} \\ K_{ff}^{\text{post}} &= K_{ff} - K_{f\mathbf{x}} K_{\mathbf{xx}}^{-1} K_{\mathbf{x}f}\end{aligned}$$

- Model parameters  $B_j$ ,  $D_j$  and  $S_j$  estimated by type II maximum likelihood,

$$\log p(\mathbf{x}) = N(\mathbf{x} | \mathbf{0}, K_{\mathbf{xx}})$$

# Implementation

## Riemann quadrature

- Implementation requires a discretised time.
- Compute the gradient and Hessian on a grid.
- Integrate them by approximate Riemann quadrature.
- We choose a uniform grid  $\{t_p\}_{p=1}^M$  so that  $\Delta = t_p - t_{p-1}$  is constant.
- The vector  $\mathbf{f} = \{f_p\}_{p=1}^M$  is the function  $f$  at the grid points.

$$I(t) = \int_0^t f(u) \exp(D_j u) du$$

$$I(t) \approx \sum_{p=1}^M f(t_p) \exp(D_j t_p) \Delta$$



# Log Likelihood

## Functional Gradient

- Given noise-corrupted data  $y_j(t_i)$  the log-likelihood is

$$\log p(Y|f, \theta_j) = -\frac{1}{2} \sum_{i=1}^T \sum_{j=1}^N \left[ \frac{(x_j(t_i) - y_j(t_i))^2}{\sigma_{ji}^2} - \log(\sigma_{ji}^2) \right] - \frac{NT}{2} \log(2\pi)$$

- The functional derivative of the log-likelihood wrt  $f$  is

$$\frac{\delta \log p(Y|f)}{\delta f(t)} = - \sum_{i=1}^T \Theta(t_i - t) \sum_{j=1}^N \frac{(x_j(t_i) - y_j(t_i))}{\sigma_{ji}^2} g'(f(t)) e^{-D_j(t_i - t)}$$

$\Theta(x)$  — Heaviside step function.

# Log Likelihood

## Functional Hessian

- Given noise-corrupted data  $y_j(t_i)$  the log-likelihood is

$$\log p(Y|f, \theta_j) = -\frac{1}{2} \sum_{i=1}^T \sum_{j=1}^N \left[ \frac{(x_j(t_i) - y_j(t_i))^2}{\sigma_{ji}^2} - \log(\sigma_{ji}^2) \right] - \frac{NT}{2} \log(2\pi)$$

- The negative Hessian of the log-likelihood wrt  $f$  is

$$\begin{aligned} w(t, t') &= \sum_{i=1}^T \Theta(t_i - t) \delta(t - t') \sum_{j=1}^N \frac{(x_j(t_i) - y_j(t_i))}{\sigma_{ji}^2} g''(f(t)) e^{-D_j(t_i - t)} \\ &\quad + \sum_{i=1}^T \Theta(t_i - t) \Theta(t_i - t') \sum_{j=1}^N \sigma_{ji}^{-2} g'(f(t)) g'(f(t')) e^{-D_j(2t_i - t - t')} \end{aligned}$$

$$g'(f) = \partial g / \partial f \text{ and } g''(f) = \partial^2 g / \partial f^2.$$

# Implementation II

## Combine with Prior

- Combine these with prior to compute gradient and Hessian of log posterior  $\Psi(\mathbf{f}) = \log p(Y|\mathbf{f}) + \log p(\mathbf{f})$  [see Rasmussen and Williams, 2006, chapter 3]

$$\begin{aligned}\frac{\partial \Psi(\mathbf{f})}{\partial \mathbf{f}} &= \frac{\partial \log p(Y|\mathbf{f})}{\partial \mathbf{f}} - K^{-1}\mathbf{f} \\ \frac{\partial^2 \Psi(\mathbf{f})}{\partial \mathbf{f}^2} &= -(W + K^{-1})\end{aligned}\tag{7}$$

$K$  prior covariance evaluated at the grid points.

- Use to find a MAP solution via,  $\hat{\mathbf{f}}$ , using Newton's algorithm.
- The Laplace approximation is then

$$\log p(Y) \simeq \log p(Y|\hat{\mathbf{f}}) - \frac{1}{2}\hat{\mathbf{f}}^T K^{-1}\hat{\mathbf{f}} - \frac{1}{2} \log |I + KW|.\tag{8}$$

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